



Bivariate Gaussian

Lecture 8

EE 640
Stochastic Systems



Outline

- Linearity of Expectation
- Correlation Coefficient
- Joint (Bivariate)Gaussian pdf
- Conditional Statistics



Linearity of Expectation

$$E\{aX + bY\} = aE\{X\} + bE\{Y\}$$

$$E\left\{\sum_{i=1}^N a_i X_i\right\} = \sum_{i=1}^N a_i E\{X_i\}$$

Conditional Expectation

$$\begin{aligned} E\{E\{g(X, Y) | X\}\} &= \int_{-\infty}^{\infty} \left[\int_{-\infty}^{\infty} g(x, y) f_{Y|X}(y | x) dy \right] f_X(x) dx \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{Y|X}(y | x) f_X(x) dx dy = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{XY}(x, y) dx dy \end{aligned}$$



Correlation Coefficient

$$\rho_{xy} = \frac{\sigma_{xy}}{\sqrt{\sigma_x^2 \sigma_y^2}} = \frac{\sigma_{xy}}{\sigma_x \sigma_y}$$

$$\mu_x = E\{X\}, \quad \mu_y = E\{Y\}$$

$$\sigma_x^2 = E\{(X - \mu_x)^2\}$$

$$\sigma_y^2 = E\{(Y - \mu_y)^2\}$$

$$\text{Uncorrelated } E\{XY\} = E\{X\}E\{Y\} \quad \sigma_{xy} = E\{(X - \mu_x)(Y - \mu_y)\}$$

$$\text{Orthogonal } E\{XY\} = 0$$

If X independent of Y, then uncorrelated

If uncorrelated and $\mu_x = \mu_y = 0$, then orthogonal

If $\rho = 0$, then uncorrelated

If $\rho = 1$, then $\sigma_{xy} = \sigma_x \sigma_y = \sigma^2$



Joint Gaussian pdf (Bivariate Gaussian)

$$f_{XY}(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_x)^2}{\sigma_x^2} - 2\rho\frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} + \frac{(y-\mu_y)^2}{\sigma_y^2}\right]\right\}$$

$$\rho = 0$$

$$\begin{aligned} f_{XY}(x, y) &= \frac{1}{2\pi\sigma_x\sigma_y} \exp\left\{-\frac{1}{2}\left[\frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2}\right]\right\} \\ &= \frac{1}{\sqrt{2\pi}\sigma_x} \exp\left\{-\frac{(x-\mu_x)^2}{2\sigma_x^2}\right\} \frac{1}{\sqrt{2\pi}\sigma_y} \exp\left\{-\frac{(y-\mu_y)^2}{2\sigma_y^2}\right\} \\ &= f_X(x)f_Y(y) \end{aligned}$$



Example 2.33

$$X \sim N(0, \sigma_x^2) \quad Y \sim N(0, \sigma_y^2)$$

$$f_{XY}(x, y) = f_X(x)f_Y(y)$$

$$Z = \frac{1}{2}(X + Y) \quad W = \frac{1}{2}(X - Y)$$

(1) $f_{ZW}(z, w) = ?$ (2) find marginals (3) Are Z and W independent?

Sum or difference of Gaussians is another Gaussian



Example Cont'

$$(2) E\{Z\} = \frac{1}{2} (E\{X\} + E\{Y\}) = 0 \quad E\{W\} = \frac{1}{2} (E\{X\} - E\{Y\}) = 0$$

$$\begin{aligned}\sigma_z^2 = E\{Z^2\} &= \frac{1}{4} E\{X^2 + 2XY + Y^2\} \\ &= \frac{1}{4} (E\{X^2\} + 2E\{XY\} + E\{Y^2\}) \\ &= [\sigma_x^2 + \sigma_y^2 + 2E\{XY\}]/4\end{aligned}$$

$$\begin{aligned}\sigma_w^2 = E\{W^2\} &= \frac{1}{4} E\{X^2 - 2XY + Y^2\} \\ &= \frac{1}{4} (E\{X^2\} - 2E\{XY\} + E\{Y^2\}) \\ &= [\sigma_x^2 + \sigma_y^2 - 2E\{XY\}]/4\end{aligned}$$

$$f_z(z) = \frac{1}{\sqrt{2\pi}\sigma_z} e^{-\frac{z^2}{2\sigma_z^2}} \quad f_w(w) = \frac{1}{\sqrt{2\pi}\sigma_w} e^{-\frac{w^2}{2\sigma_w^2}}$$



Example Cont'

$$(3) E\{ZW\} = \frac{1}{4} E\{X^2 - Y^2\} = [\sigma_x^2 - \sigma_y^2]/4$$

$$\rho_{zw} = \frac{\sigma_{zw}}{\sigma_z\sigma_w} = \frac{\frac{\sigma_x^2 - \sigma_y^2}{4}}{\frac{\sigma_x^2 + \sigma_y^2}{4}} = \frac{\sigma_x^2 - \sigma_y^2}{\sigma_x^2 + \sigma_y^2}$$

If $\sigma_x^2 = \sigma_y^2$, then Z and W are independent, else not independent



Example Cont'

$$(3) \quad f_{XY}(x, y) = \frac{1}{2\pi\sigma^2\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)\sigma^2} [z^2 - 2\rho zw + w^2]\right\}$$

Where $\sigma^2 = \sigma_z^2 = \sigma_w^2$, $\rho = \rho_{zw}$



Conditional Statistics

$$F_X(x) = \sum_{i=1}^n F_X(x | A_i) P(A_i)$$

$$\frac{\partial F_X(x)}{\partial x} = f_X(x) = \sum_{i=1}^n \frac{\partial F(x | A_i)}{\partial x} P(A_i)$$

Where $\frac{\partial F(x | A_i)}{\partial x} = \frac{\partial}{\partial x} \int_{-\infty}^x f(\lambda | A_i) d\lambda$

$$f_X(x) = \sum_{i=1}^n f_X(x | A_i) P(A_i)$$